

# Appendix A

## List of Distributions

Here we list common statistical distributions used throughout the book. The often used indicator symbol  $1_{\{\cdot\}}$  and gamma function  $\Gamma(\alpha)$  are defined as follows.

**Definition A.1** The indicator symbol is defined as

$$1_{\{\cdot\}} = \begin{cases} 1, & \text{if condition in } \{\cdot\} \text{ is true,} \\ 0, & \text{otherwise.} \end{cases} \tag{A.1}$$

**Definition A.2** The standard gamma function is defined as

$$\Gamma(\alpha) = \int_0^\infty t^{\alpha-1} e^{-t} dt, \quad \alpha > 0. \tag{A.2}$$

### A.1 Discrete Distributions

#### A.1.1 Poisson Distribution, $Poisson(\lambda)$

A Poisson distribution function is denoted as  $Poisson(\lambda)$ . The random variable  $N$  has a Poisson distribution, denoted  $N \sim Poisson(\lambda)$ , if its probability mass function is

$$p(k) = \Pr[N = k] = \frac{\lambda^k}{k!} e^{-\lambda}, \quad \lambda > 0 \tag{A.3}$$

for all  $k \in \{0, 1, 2, \dots\}$ . Expectation, variance and variational coefficient of a random variable  $N \sim Poisson(\lambda)$  are

$$E[N] = \lambda, \quad \text{Var}[N] = \lambda, \quad \text{Vco}[N] = \frac{1}{\sqrt{\lambda}}. \tag{A.4}$$

### A.1.2 Binomial Distribution, $Bin(n, p)$

The binomial distribution function is denoted as  $Bin(n, p)$ . The random variable  $N$  has a binomial distribution, denoted  $N \sim Bin(n, p)$ , if its probability mass function is

$$p(k) = \Pr[N = k] = \binom{n}{k} p^k (1-p)^{n-k}, \quad p \in (0, 1), \quad n \in 1, 2, \dots \quad (\text{A.5})$$

for all  $k \in \{0, 1, 2, \dots, n\}$ . Expectation, variance and variational coefficient of a random variable  $N \sim Bin(n, p)$  are

$$E[N] = np, \quad \text{Var}[N] = np(1-p), \quad \text{Vco}[N] = \sqrt{\frac{1-p}{np}}. \quad (\text{A.6})$$

*Remark A.1*  $N$  is the number of successes in  $n$  independent trials, where  $p$  is the probability of a success in each trial.

### A.1.3 Negative Binomial Distribution, $NegBin(r, p)$

A negative binomial distribution function is denoted as  $NegBin(r, p)$ . The random variable  $N$  has a negative binomial distribution, denoted  $N \sim NegBin(r, p)$ , if its probability mass function is

$$p(k) = \Pr[N = k] = \binom{r+k-1}{k} p^r (1-p)^k, \quad p \in (0, 1), \quad r \in (0, \infty) \quad (\text{A.7})$$

for all  $k \in \{0, 1, 2, \dots\}$ . Here, the generalised binomial coefficient is

$$\binom{r+k-1}{k} = \frac{\Gamma(k+r)}{k! \Gamma(r)}, \quad (\text{A.8})$$

where  $\Gamma(r)$  is the gamma function.

Expectation, variance and variational coefficient of a random variable  $N \sim NegBin(r, p)$  are

$$E[N] = \frac{r(1-p)}{p}, \quad \text{Var}[N] = \frac{r(1-p)}{p^2}, \quad \text{Vco}[N] = \frac{1}{\sqrt{r(1-p)}}. \quad (\text{A.9})$$

*Remark A.2* If  $r$  is a positive integer,  $N$  is the number of failures in a sequence of independent trials until  $r$  successes, where  $p$  is the probability of a success in each trial.

## A.2 Continuous Distributions

### A.2.1 Uniform Distribution, $\mathcal{U}(a, b)$

A uniform distribution function is denoted as  $\mathcal{U}(a, b)$ . The random variable  $X$  has a uniform distribution, denoted  $X \sim \mathcal{U}(a, b)$ , if its probability density function is

$$f(x) = \frac{1}{b-a}, \quad a < b \quad (\text{A.10})$$

for  $x \in [a, b]$ . Expectation, variance and variational coefficient of a random variable  $X \sim \mathcal{U}(a, b)$  are

$$E[X] = \frac{a+b}{2}, \quad \text{Var}[X] = \frac{(b-a)^2}{12}, \quad \text{Vco}[X] = \frac{b-a}{\sqrt{3}(a+b)}. \quad (\text{A.11})$$

### A.2.2 Normal (Gaussian) Distribution, $\mathcal{N}(\mu, \sigma)$

A normal (Gaussian) distribution function is denoted as  $\mathcal{N}(\mu, \sigma)$ . The random variable  $X$  has a normal distribution, denoted  $X \sim \mathcal{N}(\mu, \sigma)$ , if its probability density function is

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right), \quad \sigma^2 > 0, \quad \mu \in \mathbb{R} \quad (\text{A.12})$$

for all  $x \in \mathbb{R}$ . Expectation, variance and variational coefficient of a random variable  $X \sim \mathcal{N}(\mu, \sigma)$  are

$$E[X] = \mu, \quad \text{Var}[X] = \sigma^2, \quad \text{Vco}[X] = \sigma/\mu. \quad (\text{A.13})$$

### A.2.3 Lognormal Distribution, $\mathcal{LN}(\mu, \sigma)$

A lognormal distribution function is denoted as  $\mathcal{LN}(\mu, \sigma)$ . The random variable  $X$  has a lognormal distribution, denoted  $X \sim \mathcal{LN}(\mu, \sigma)$ , if its probability density function is

$$f(x) = \frac{1}{x\sqrt{2\pi}\sigma} \exp\left(-\frac{(\ln(x)-\mu)^2}{2\sigma^2}\right), \quad \sigma^2 > 0, \quad \mu \in \mathbb{R} \quad (\text{A.14})$$

for  $x > 0$ . Expectation, variance and variational coefficient of a random variable  $X \sim \mathcal{LN}(\mu, \sigma)$  are

$$E[X] = e^{\mu + \frac{1}{2}\sigma^2}, \quad \text{Var}[X] = e^{2\mu + \sigma^2}(e^{\sigma^2} - 1), \quad \text{Vco}[X] = \sqrt{e^{\sigma^2} - 1}. \quad (\text{A.15})$$

### A.2.4 *t Distribution, $\mathcal{T}(v, \mu, \sigma^2)$*

A *t* distribution function is denoted as  $\mathcal{T}(v, \mu, \sigma^2)$ . The random variable  $X$  has a *t* distribution, denoted  $X \sim \mathcal{T}(v, \mu, \sigma^2)$ , if its probability density function is

$$f(x) = \frac{\Gamma((v+1)/2)}{\Gamma(v/2)} \frac{1}{\sqrt{v\pi}} \left(1 + \frac{(x-\mu)^2}{v\sigma^2}\right)^{-(v+1)/2} \quad (\text{A.16})$$

for  $\sigma^2 > 0$ ,  $\mu \in \mathbb{R}$ ,  $v = 1, 2, \dots$  and all  $x \in \mathbb{R}$ . Expectation, variance and variational coefficient of a random variable  $X \sim \mathcal{T}(v, \mu, \sigma^2)$  are

$$\begin{aligned} E[X] &= \mu \text{ if } v > 1, \\ \text{Var}[X] &= \sigma^2 \frac{v}{v-2} \text{ if } v > 2, \\ \text{Vco}[X] &= \frac{\sigma}{\mu} \sqrt{\frac{v}{v-2}} \text{ if } v > 2. \end{aligned} \quad (\text{A.17})$$

### A.2.5 *Gamma Distribution, $\text{Gamma}(\alpha, \beta)$*

A gamma distribution function is denoted as  $\text{Gamma}(\alpha, \beta)$ . The random variable  $X$  has a gamma distribution, denoted as  $X \sim \text{Gamma}(\alpha, \beta)$ , if its probability density function is

$$f(x) = \frac{x^{\alpha-1}}{\Gamma(\alpha)\beta^\alpha} \exp(-x/\beta), \quad \alpha > 0, \quad \beta > 0 \quad (\text{A.18})$$

for  $x > 0$ . Expectation, variance and variational coefficient of a random variable  $X \sim \text{Gamma}(\alpha, \beta)$  are

$$E[X] = \alpha\beta, \quad \text{Var}[X] = \alpha\beta^2, \quad \text{Vco}[X] = 1/\sqrt{\alpha}. \quad (\text{A.19})$$

### A.2.6 *Weibull Distribution, $\text{Weibull}(\alpha, \beta)$*

A Weibull distribution function is denoted as  $\text{Weibull}(\alpha, \beta)$ . The random variable  $X$  has a Weibull distribution, denoted as  $X \sim \text{Weibull}(\alpha, \beta)$ , if its probability density function is

$$f(x) = \frac{\alpha}{\beta^\alpha} x^{\alpha-1} \exp(-(x/\beta)^\alpha), \quad \alpha > 0, \quad \beta > 0 \quad (\text{A.20})$$

for  $x > 0$ . The corresponding distribution function is

$$F(x) = 1 - \exp(-(x/\beta)^\alpha), \quad \alpha > 0, \beta > 0. \tag{A.21}$$

Expectation and variance of a random variable  $X \sim Weibull(\alpha, \beta)$  are

$$E[X] = \beta\Gamma(1 + 1/\alpha), \quad \text{Var}[X] = \beta^2 \left( \Gamma(1 + 2/\alpha) - (\Gamma(1 + 1/\alpha))^2 \right).$$

### A.2.7 Pareto Distribution (One-Parameter), $Pareto(\xi, x_0)$

A one-parameter Pareto distribution function is denoted as  $Pareto(\xi, x_0)$ . The random variable  $X$  has a Pareto distribution, denoted as  $X \sim Pareto(\xi, x_0)$ , if its distribution function is

$$F(x) = 1 - \left( \frac{x}{x_0} \right)^{-\xi}, \quad x \geq x_0, \tag{A.22}$$

where  $x_0 > 0$  and  $\xi > 0$ . The support starts at  $x_0$ , which is typically known and not considered as a parameter. Therefore the distribution is referred to as a single parameter Pareto. The corresponding probability density function is

$$f(x) = \frac{\xi}{x_0} \left( \frac{x}{x_0} \right)^{-\xi-1}. \tag{A.23}$$

Expectation, variance and variational coefficient of  $X \sim Pareto(\xi, x_0)$  are

$$\begin{aligned} E[X] &= x_0 \frac{\xi}{\xi - 1} \quad \text{if } \xi > 1, \\ \text{Var}[X^2] &= x_0^2 \frac{\xi}{(\xi - 1)^2(\xi - 2)} \quad \text{if } \xi > 2, \\ \text{Vco}[X] &= \frac{1}{\sqrt{\xi(\xi - 2)}} \quad \text{if } \xi > 2. \end{aligned}$$

### A.2.8 Pareto Distribution (Two-Parameter), $Pareto_2(\alpha, \beta)$

A two-parameter Pareto distribution function is denoted as  $Pareto_2(\alpha, \beta)$ . The random variable  $X$  has a Pareto distribution, denoted as  $X \sim Pareto_2(\alpha, \beta)$ , if its distribution function is

$$F(x) = 1 - \left( 1 + \frac{x}{\beta} \right)^{-\alpha}, \quad x \geq 0, \tag{A.24}$$

where  $\alpha > 0$  and  $\beta > 0$ . The corresponding probability density function is

$$f(x) = \frac{\alpha\beta^\alpha}{(x + \beta)^{\alpha+1}}. \quad (\text{A.25})$$

The moments of a random variable  $X \sim \text{Pareto}_2(\alpha, \beta)$  are

$$E[X^k] = \frac{\beta^k k!}{\prod_{i=1}^k (\alpha - i)}; \quad \alpha > k.$$

### A.2.9 Generalised Pareto Distribution, $GPD(\xi, \beta)$

A GPD distribution function is denoted as  $GPD(\xi, \beta)$ . The random variable  $X$  has a GPD distribution, denoted as  $X \sim GPD(\xi, \beta)$ , if its distribution function is

$$H_{\xi, \beta}(x) = \begin{cases} 1 - (1 + \xi x/\beta)^{-1/\xi}, & \xi \neq 0, \\ 1 - \exp(-x/\beta), & \xi = 0, \end{cases} \quad (\text{A.26})$$

where  $x \geq 0$  when  $\xi \geq 0$  and  $0 \leq x \leq -\beta/\xi$  when  $\xi < 0$ . The corresponding probability density function is

$$h(x) = \begin{cases} \frac{1}{\beta} (1 + \xi x/\beta)^{-\frac{1}{\xi}-1}, & \xi \neq 0, \\ \frac{1}{\beta} \exp(-x/\beta), & \xi = 0. \end{cases} \quad (\text{A.27})$$

Expectation, variance and variational coefficient of  $X \sim GPD(\xi, \beta)$ ,  $\xi \geq 0$ , are

$$\begin{aligned} E[X^n] &= \frac{\beta^n n!}{\prod_{k=1}^n (1 - k\xi)}, \quad \xi < \frac{1}{n}; \quad E[X] = \frac{\beta}{1 - \xi}, \quad \xi < 1; \\ \text{Var}[X^2] &= \frac{\beta^2}{(1 - \xi)^2(1 - 2\xi)}, \quad \text{Vco}[X] = \frac{1}{\sqrt{1 - 2\xi}}, \quad \xi < \frac{1}{2}. \end{aligned} \quad (\text{A.28})$$

### A.2.10 Beta Distribution, $Beta(\alpha, \beta)$

A beta distribution function is denoted as  $Beta(\alpha, \beta)$ . The random variable  $X$  has a beta distribution, denoted as  $X \sim Beta(\alpha, \beta)$ , if its probability density function is

$$f(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}, \quad 0 \leq x \leq 1, \quad (\text{A.29})$$

for  $\alpha > 0$  and  $\beta > 0$ . Expectation, variance and variational coefficient of a random variable  $X \sim Beta(\alpha, \beta)$  are

$$E[X] = \frac{\alpha}{\alpha + \beta}, \quad \text{Var}[X] = \frac{\alpha\beta}{(\alpha + \beta)^2(1 + \alpha + \beta)}, \quad \text{Vco}[X] = \sqrt{\frac{\beta}{\alpha(1 + \alpha + \beta)}}.$$

### A.2.11 Generalised Inverse Gaussian Distribution, $GIG(\omega, \phi, \nu)$

A GIG distribution function is denoted as  $GIG(\omega, \phi, \nu)$ . The random variable  $X$  has a GIG distribution, denoted as  $X \sim GIG(\omega, \phi, \nu)$ , if its probability density function is

$$f(x) = \frac{(\omega/\phi)^{(\nu+1)/2}}{2K_{\nu+1}(2\sqrt{\omega\phi})} x^\nu e^{-x\omega - x^{-1}\phi}, \quad x > 0, \quad (\text{A.30})$$

where  $\phi > 0, \omega \geq 0$  if  $\nu < -1$ ;  $\phi > 0, \omega > 0$  if  $\nu = -1$ ;  $\phi \geq 0, \omega > 0$  if  $\nu > -1$ ; and

$$K_{\nu+1}(z) = \frac{1}{2} \int_0^\infty u^\nu e^{-z(u+1/u)/2} du.$$

$K_\nu(z)$  is called a modified Bessel function of the third kind; see for instance Abramowitz and Stegun ([3], p. 375).

The moments of a random variable  $X \sim GIG(\omega, \phi, \nu)$  are not available in a closed form through elementary functions but can be expressed in terms of Bessel functions:

$$E[X^\alpha] = \left(\frac{\phi}{\omega}\right)^{\alpha/2} \frac{K_{\nu+1+\alpha}(2\sqrt{\omega\phi})}{K_{\nu+1}(2\sqrt{\omega\phi})}, \quad \alpha \geq 1, \quad \phi > 0, \quad \omega > 0.$$

Often, using notation  $R_\nu(z) = K_{\nu+1}(z)/K_\nu(z)$ , it is written as

$$E[X^\alpha] = \left(\frac{\phi}{\omega}\right)^{\alpha/2} \prod_{k=1}^\alpha R_{\nu+k}(2\sqrt{\omega\phi}), \quad \alpha = 1, 2, \dots$$

The mode is easily calculated from  $\frac{\partial}{\partial x} x^\nu e^{-(\omega x + \phi/x)} = 0$  as

$$\text{mode}[X] = \frac{1}{2\omega} \left( \nu + \sqrt{\nu^2 + 4\omega\phi} \right),$$

that differs only slightly from the expected value for large  $\nu$ , i.e.

$$\text{mode}[X] \rightarrow E[X] \quad \text{for } \nu \rightarrow \infty.$$

### A.2.12 $d$ -variate Normal Distribution, $\mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

A  $d$ -variate normal distribution function is denoted as  $\mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ , where  $\boldsymbol{\mu} = (\mu_1, \dots, \mu_d)' \in \mathbb{R}^d$  and  $\boldsymbol{\Sigma}$  is a positive definite matrix ( $d \times d$ ). The corresponding probability density function is

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} \sqrt{\det \boldsymbol{\Sigma}}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right), \quad \mathbf{x} \in \mathbb{R}^d, \quad (\text{A.31})$$

where  $\boldsymbol{\Sigma}^{-1}$  is the inverse of the matrix  $\boldsymbol{\Sigma}$ . Expectations and covariances of a random vector  $\mathbf{X} = (X_1, \dots, X_d)' \sim \mathcal{N}_d(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  are

$$\mathbb{E}[X_i] = \mu_i, \quad \text{Cov}[X_i, X_j] = \Sigma_{i,j}, \quad i, j = 1, \dots, d. \quad (\text{A.32})$$

### A.2.13 $d$ -variate $t$ -Distribution, $\mathcal{T}_d(\nu, \boldsymbol{\mu}, \boldsymbol{\Sigma})$

A  $d$ -variate  $t$ -distribution function with  $\nu$  degrees of freedom is denoted as  $\mathcal{T}_d(\nu, \boldsymbol{\mu}, \boldsymbol{\Sigma})$ , where  $\nu > 0$ ,  $\boldsymbol{\mu} = (\mu_1, \dots, \mu_d)' \in \mathbb{R}^d$  is a location vector and  $\boldsymbol{\Sigma}$  is a positive definite matrix ( $d \times d$ ). The corresponding probability density function is

$$f(\mathbf{x}) = \frac{\Gamma\left(\frac{\nu+d}{2}\right)}{(\nu\pi)^{d/2} \Gamma\left(\frac{\nu}{2}\right) \sqrt{\det \boldsymbol{\Sigma}}} \left(1 + \frac{(\mathbf{x} - \boldsymbol{\mu})' \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})}{\nu}\right)^{-\frac{\nu+d}{2}}, \quad (\text{A.33})$$

where  $\mathbf{x} \in \mathbb{R}^d$  and  $\boldsymbol{\Sigma}^{-1}$  is the inverse of the matrix  $\boldsymbol{\Sigma}$ . Expectations and covariances of a random vector  $\mathbf{X} = (X_1, \dots, X_d)' \sim \mathcal{T}_d(\nu, \boldsymbol{\mu}, \boldsymbol{\Sigma})$  are

$$\begin{aligned} \mathbb{E}[X_i] &= \mu_i, \quad \text{if } \nu > 1, \quad i = 1, \dots, d; \\ \text{Cov}[X_i, X_j] &= \nu \Sigma_{i,j} / (\nu - 2), \quad \text{if } \nu > 2, \quad i, j = 1, \dots, d. \end{aligned} \quad (\text{A.34})$$



## Appendix B

### Selected Simulation Algorithms

#### B.1 Simulation from GIG Distribution

To generate realisations of a random variable  $X \sim \text{GIG}(\omega, \phi, \nu)$  with  $\omega, \phi > 0$ , a special algorithm is required because we cannot invert the distribution function in closed form. The following algorithm can be found in Dagpunar [67]:

##### Algorithm B.1 (Simulation from GIG distribution)

1.  $\alpha = \sqrt{\omega/\phi}$ ;  $\beta = 2\sqrt{\omega\phi}$ ,  

$$m = \frac{1}{\beta} \left( \nu + \sqrt{\nu^2 + \beta^2} \right),$$

$$g(y) = \frac{1}{2}\beta y^3 - y^2 \left( \frac{1}{2}\beta m + \nu + 2 \right) + y \left( \nu m - \frac{\beta}{2} \right) + \frac{1}{2}\beta m.$$
2. Set  $y_0 = m$ ,  
 While  $g(y_0) \leq 0$  do  $y_0 = 2y_0$ ,  
 $y_+$ : root of  $g$  in the interval  $(m, y_0)$ ,  
 $y_-$ : root of  $g$  in the interval  $(0, m)$ .
3.  $a = (y_+ - m) \left( \frac{y_+}{m} \right)^{\nu/2} \exp \left( -\frac{\beta}{4} \left( y_+ + \frac{1}{y_+} - m - \frac{1}{m} \right) \right)$ ,  
 $b = (y_- - m) \left( \frac{y_-}{m} \right)^{\nu/2} \exp \left( -\frac{\beta}{4} \left( y_- + \frac{1}{y_-} - m - \frac{1}{m} \right) \right)$ ,  
 $c = -\frac{\beta}{4} \left( m + \frac{1}{m} \right) + \frac{\nu}{2} \ln(m).$
4. Repeat  $U, V \sim \mathcal{U}(0, 1)$ ,  $Y = m + a\frac{U}{V} + b\frac{1-V}{U}$ ,  
 until  $Y > 0$  and  $-\ln U \geq -\frac{\nu}{2} \ln Y + \frac{1}{4}\beta \left( Y + \frac{1}{Y} \right) + c$ ,  
 Then  $X = \frac{Y}{\alpha}$  is  $\text{GIG}(\omega, \phi, \nu)$ .

To generate a sequence of  $n$  realisations from a GIG random variable, step 4 is repeated  $n$  times.

## B.2 Simulation from $\alpha$ -stable Distribution

To generate realisations of a random variable  $X \sim \alpha\text{Stable}(\alpha, \beta, \sigma, \mu)$ , defined by (6.56), a special algorithm is required because the density of  $\alpha$ -stable distribution is not available in closed form. An elegant and efficient solution was proposed in Chambers, Mallows and Stuck [50]; also see Nolan [176].

### Algorithm B.2 (Simulation from $\alpha$ -stable distribution)

1. Simulate  $W$  from the exponential distribution with mean = 1.
2. Simulate  $U$  from  $\mathcal{U}\left(-\frac{\pi}{2}, \frac{\pi}{2}\right)$ .
3. Calculate

$$Z = \begin{cases} S_{\alpha,\beta} \frac{\sin(\alpha(U+B_{\alpha,\beta}))}{(\cos U)^{1/\alpha}} \left( \frac{\cos(U-\alpha(U+B_{\alpha,\beta}))}{W} \right)^{-1+\frac{1}{\alpha}}, & \alpha \neq 1, \\ \frac{2}{\pi} \left( \left( \frac{\pi}{2} + \beta U \right) \tan U - \beta \ln \left( \frac{\pi W \cos U}{\pi + 2\beta U} \right) \right), & \alpha = 1, \end{cases}$$

where

$$S_{\alpha,\beta} = (1 + \beta^2 \tan^2(\pi\alpha/2))^{1/(2\alpha)},$$

$$B_{\alpha,\beta} = \frac{1}{\alpha} \arctan(\beta \tan(\pi\alpha/2)).$$

The obtained  $Z$  is a sample from  $\alpha\text{Stable}(\alpha, \beta, 1, 0)$ .

4. Then,

$$X = \begin{cases} \mu + \sigma Z, & \alpha \neq 1, \\ \mu + \sigma Z + \frac{2}{\pi} \beta \sigma \ln \sigma, & \alpha = 1, \end{cases}$$

is a sample from  $\alpha\text{Stable}(\alpha, \beta, \sigma, \mu)$ .

Note that there are different parameterisations of the  $\alpha$ -stable distribution. The algorithm above is for representation (6.56).

# Solutions for Selected Problems

## Problems of Chapter 2

**2.1** The likelihood function for independent data  $\mathbf{N} = \{N_1, N_2, \dots, N_M\}$  from *Poisson*( $\lambda$ ) is

$$\ell_{\mathbf{n}}(\lambda) = \prod_{i=1}^M e^{-\lambda} \frac{\lambda^{n_i}}{n_i!},$$
$$\ln \ell_{\mathbf{n}}(\lambda) = -\lambda M + \ln \lambda \sum_{i=1}^M n_i - \sum_{i=1}^M \ln(n_i!).$$

The MLE  $\hat{\Lambda}$  maximising the log-likelihood function  $\ln \ell_{\mathbf{N}}(\lambda)$  is

$$\hat{\Lambda} = \frac{1}{M} \sum_{i=1}^M N_i.$$

Using the properties of the Poisson distribution,  $E[N_i] = \text{Var}[N_i] = \lambda$ , it is easy to get

$$E[\hat{\Lambda}] = \frac{1}{M} \sum_{i=1}^M E[N_i] = \lambda;$$
$$\text{Var}[\hat{\Lambda}] = \frac{1}{M^2} \sum_{i=1}^M \text{Var}[N_i] = \frac{\lambda}{M}.$$

To estimate the variance of  $\hat{\Lambda}$  using a normal approximation, find the information matrix

$$I(\lambda) = -\frac{1}{M} E \left[ \frac{\partial^2 \ln \ell_{\mathbf{N}}(\lambda)}{\partial \lambda^2} \right] = \frac{1}{M \lambda^2} E \left[ \sum_{i=1}^M N_i \right] = \frac{1}{\lambda}.$$

Thus, using asymptotic normal distribution approximation,

$$\text{Var}[\widehat{\Lambda}] \approx \text{I}^{-1}(\lambda)/M = \lambda/M.$$

In both cases the variance depends on unknown true parameter  $\lambda$  that can be estimated, for a given realisation  $\mathbf{n}$ , as  $\widehat{\lambda}$ .

**2.4** Consider

$$L(\mathbf{u}) = u_1 L_1 + \cdots + u_J L_J,$$

where  $\mathbf{u} \in \mathbb{R}^J$  and set

$$\phi_{\mathbf{u}}(t) = \varrho[tL(\mathbf{u})], \quad t > 0.$$

Then using homogeneity property  $\varrho[tL] = t\varrho[L]$ ,

$$\frac{d\phi_{\mathbf{u}}(t)}{dt} = \varrho[L(\mathbf{u})].$$

From another side

$$\frac{d\phi_{\mathbf{u}}(t)}{dt} = \sum_{j=1}^J \frac{\varrho[L(\mathbf{x})]}{\partial x_j} \Big|_{\mathbf{x}=t\mathbf{u}} u_j = \sum_{j=1}^J \frac{\varrho[L(\mathbf{u})]}{\partial u_j} u_j,$$

where to get the last equality we used homogeneity property. Thus

$$\varrho[L(\mathbf{1})] = \sum_{j=1}^J \frac{\varrho[L_1 + \cdots + L_j + hL_j]}{\partial h} \Big|_{h=0}$$

completes the proof.

**2.5** The sum of risks is gamma distributed:

$$Z_1 + Z_2 + Z_3 \sim \text{Gamma}(\alpha_1 + \alpha_2 + \alpha_3, \beta).$$

Thus  $\text{VaR}_{0.999}[Z_i] = F_G^{-1}(0.999|\alpha_i, \beta)$  and

$$\text{VaR}_{0.999}[Z_1 + Z_2 + Z_3] = F_G^{-1}(0.999|\alpha_1 + \alpha_2 + \alpha_3, \beta),$$

where  $F_G^{-1}(\cdot|\alpha, \beta)$  is the inverse of the  $\text{Gamma}(\alpha, \beta)$ . Using, for example, MS Excel spreadsheet function  $\text{GAMMAINV}(\cdot)$ , find

$$\begin{aligned} \text{VaR}_{0.999}[Z_1] &\approx 5.414, & \text{VaR}_{0.999}[Z_2] &\approx 6.908, \\ \text{VaR}_{0.999}[Z_3] &\approx 8.133, & \text{VaR}_{0.999}[Z_1 + Z_2 + Z_3] &\approx 11.229. \end{aligned}$$

The sum of VaRs is  $\text{VaR}_{0,999}[Z_1] + \text{VaR}_{0,999}[Z_2] + \text{VaR}_{0,999}[Z_3] \approx 20.455$  and thus the diversification is  $\approx 45\%$ .

### Problems of Chapter 3

**3.1** By definition of the expected shortfall we have

$$\begin{aligned} E[Z|Z > L] &= \frac{1}{1 - H(L)} \int_L^\infty zh(z)dz \\ &= \frac{E[Z]}{1 - H(L)} - \frac{1}{1 - H(L)} \int_0^L zh(z)dz. \end{aligned}$$

Substituting  $h(z)$  calculated via characteristic function (3.11) and changing variable  $x = t \times L$ , we obtain

$$\begin{aligned} \int_0^L zh(z)dz &= \frac{2}{\pi} \int_0^L z \int_0^\infty \text{Re}[\chi(t)] \cos(tz) dt dz \\ &= \frac{2L}{\pi} \int_0^\infty \text{Re}[\chi(x/L)] \left[ \frac{\sin(x)}{x} - \frac{1 - \cos(x)}{x^2} \right] dx. \end{aligned}$$

Recognizing that the term involving  $\sin(x)/x$  corresponds to  $H(L)$ , we obtain

$$E[Z|Z > L] = \frac{1}{1 - H(L)} \left[ E[Z] - H(L)L + \frac{2L}{\pi} \int_0^\infty \text{Re}[\chi(x/L)] \frac{1 - \cos x}{x^2} dx \right].$$

### Problems of Chapter 4

**4.1** The linear estimator  $\widehat{\theta}_{tot} = w_1\widehat{\theta}_1 + \dots + w_K\widehat{\theta}_K$  is unbiased, i.e.  $E[\widehat{\theta}_{tot}] = \theta$ , if  $w_1 + \dots + w_K = 1$  because  $E[\widehat{\theta}_k] = \theta$ . Minimisation of the variance

$$\text{Var}[\widehat{\theta}_{tot}] = w_1^2\sigma_1^2 + \dots + w_K^2\sigma_K^2$$

under the constraint  $w_1 + \dots + w_K = 1$  is equivalent to unconstrained minimisation of the

$$\Psi = \text{Var}[\widehat{\theta}_{tot}] - \lambda(w_1 + \dots + w_K),$$

which is a well-known method of Lagrange multipliers. Optimisation of the above requires solution of the following equations:

$$\begin{aligned}\frac{\partial \Psi}{\partial w_i} &= 2w_i \sigma_i^2 - \lambda = 0, \quad i = 1, \dots, K; \\ \frac{\partial \Psi}{\partial \lambda} &= -(w_1 + \dots + w_K) = 0.\end{aligned}$$

That gives

$$\frac{1}{2}\lambda = \left( \sum_{k=1}^K \left( 1/\sigma_k^2 \right) \right)^{-1}, \quad w_i = \frac{1/\sigma_i^2}{\sum_{k=1}^K \left( 1/\sigma_k^2 \right)}.$$

**4.2** Given  $\Theta = \theta$ , the joint density of the data at  $\mathbf{N} = \mathbf{n}$  is

$$f(\mathbf{n}|\theta) \propto \prod_{i=1}^T \theta^{n_i} (1-\theta)^{V_i - n_i}.$$

From Bayes's theorem, the posterior density of  $\theta$  is  $\pi(\theta|\mathbf{n}) \propto f(\mathbf{n}|\theta)\pi(\theta)$ , where  $\pi(\theta)$  is the prior density. Thus

$$\begin{aligned}\pi(\theta|\mathbf{n}) &\propto \theta^{\alpha-1} (1-\theta)^{\beta-1} \prod_{i=1}^T \theta^{n_i} (1-\theta)^{V_i - n_i} \\ &= \theta^{\alpha_T-1} (1-\theta)^{\beta_T-1},\end{aligned}$$

where

$$\alpha_T = \alpha + \sum_{i=1}^T n_i, \quad \beta_T = \beta + \sum_{i=1}^T V_i - \sum_{i=1}^T n_i.$$

Thus the posterior distribution of  $\Theta$  is  $Beta(\alpha_T, \beta_T)$ .

## Problems of Chapter 5

**5.1** Denote the data above  $L$  as  $\tilde{\mathbf{X}} = (\tilde{X}_1, \tilde{X}_2, \dots, \tilde{X}_k)'$ . These random variables are independent with a common density  $f(x|\xi)/(1-F(L|\xi))$ ,  $x \geq L$ , where  $f(x|\xi)$  is the density of the Pareto distribution  $F(x|\xi) = 1 - (x/a)^{-\xi}$ ,  $x \geq a > 0$ . Thus the likelihood function for given data above  $L$  is

$$\ell_{\tilde{x}}(\xi) = \prod_{i=1}^k \frac{f(\tilde{x}_i|\xi)}{1 - F(L|\xi)}.$$

Substituting the Pareto density

$$f(x|\xi) = \frac{\xi}{a} \left(\frac{x}{a}\right)^{-\xi-1}$$

gives

$$\ln \ell_{\tilde{x}}(\xi) = K\xi \ln(L/a) + K \ln(\xi/a) - (\xi + 1) \sum_{i=1}^K \ln(\tilde{x}_i/a).$$

Then, solving  $\partial \ln \ell_{\tilde{x}}(\xi)/\partial \xi = 0$ , we obtain

$$\hat{\xi}^{\text{MLE}} = \left( -\ln(L/a) + \frac{1}{K} \sum_{i=1}^K \ln(\tilde{x}_i/a) \right)^{-1}.$$

## Problems of Chapter 6

**6.1** The probability generating function of the negative binomial,  $NegBin(r, p)$ , is  $\psi(t) = (1 - (t-1)(1-p)/p)^{-r}$ . Then, using formula (6.29), we obtain that the distribution of the maximum loss over one year is

$$F_M(x) = \psi(F(x)) = \left( 1 + \frac{1-p}{p}(1-F(x)) \right)^{-r},$$

where  $F(x) = 1 - \exp(-x/\beta)$  is the severity distribution. The distribution of the maximum loss over  $m$  years is simply

$$(F_M(x))^m = \left( 1 + \frac{1-p}{p}(1-F(x)) \right)^{-r \times m}.$$

## Problems of Chapter 7

**7.1** Consider random variables  $U_1$  and  $U_2$  from the  $t$ -copula  $C_{v,\rho}^{(t)}(u_1, u_2)$ . By definition, the lower tail dependence is

$$\lambda_L = \lim_{q \rightarrow 0^+} \frac{C_{v,\rho}^{(t)}(q, q)}{q}.$$

Due to the radial symmetry of the  $t$ -copula, the upper tail dependence  $\lambda_U$  is the same as  $\lambda_L$ . Applying L'Hôpital's rule, that is, taking derivatives of the nominator and denominator,

$$\lambda_L = \lim_{q \rightarrow 0^+} \frac{dC_{v,\rho}^{(t)}(q, q)}{dq} = \lim_{q \rightarrow 0^+} \{\Pr[U_2 \leq q | U_1 = q] + \Pr[U_1 \leq q | U_2 = q]\}.$$

Let  $X_1 = F_v^{(-1)}(U_1)$  and  $X_2 = F_v^{(-1)}(U_2)$ , where  $F_v(\cdot)$  is a standard univariate  $t$ -distribution with  $v$  degrees of freedom,  $\mathcal{T}(v, 0, 1)$ . Thus  $(X_1, X_2)'$  is from a bivariate  $t$ -distribution  $\mathcal{T}_2(v, 0, \Sigma)$ , where  $\Sigma$  is a correlation matrix with off-diagonal element  $\rho$ . Then, one can calculate the conditional density of  $X_2$  given  $X_1 = x_1$ :

$$f(x_2|x_1) = \frac{f(x_1, x_2)}{f(x_1)} \propto \left(1 + \frac{v+1}{(1-\rho^2)(v+x_1^2)} \frac{(x_2 - \rho x_1)^2}{v+1}\right)^{-(v+2)/2}.$$

This can be recognised as a univariate  $t$  distribution  $\mathcal{T}(v+1, \mu, \sigma^2)$  with the mean  $\mu = \rho x_1$ ,  $\sigma^2 = \frac{(1-\rho^2)(v+x_1^2)}{v+1}$  and  $v+1$  degrees of freedom. Thus

$$\Pr[X_2 \leq x | X_1 = x] = F_{v+1} \left( \frac{(x - x\rho)\sqrt{v+1}}{\sqrt{(1-\rho^2)(v+x^2)}} \right).$$

Finally, using that  $\Pr[X_1 \leq x | X_2 = x] = \Pr[X_2 \leq x | X_1 = x]$  and taking limit  $x \rightarrow -\infty$  we get

$$\lambda = 2F_{v+1} \left( -\sqrt{\frac{(v+1)(1-\rho)}{1+\rho}} \right).$$



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